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## Linear Regression Analysis

When analyzing the Homes model, I first used regression on the model so that I could get an idea of what it looked like with multicollinearity and have something to measure the progress of subsequent models off of. This gave an overall $p$-value of $8.906\left(e^{\wedge}-8\right)$, showing that there was some multicollinearity issues. I then created and analyzed the interactions that could be taking place between the factors using the "pairs" function on RStudio. Many of the plots did not seem to have much interaction between them, even though there were some that had an obvious interaction between them, such as Age and Agesq. These two factors seem to have a hook or parabolic shape between them. Also, Bathbed had an anticipated interaction with both Bath and Bed. None of the factors, however, had an obvious and unanticipated interaction with the other factors, so they were all left in at this point in the analysis. In order to find a model that more accurately represented the price, the p-values of the factors were calculated next to see what variables are fairly unrelated to price.

Looking at the $p$-values, there were some factors that had too high of $p$-values to be used. These variables were Age (0.24784), Garage (0.22460), Adams (0.30998), Crest (0.68266), and Parker (0.38942). Out of the five dummy variable sets regarding the different elementary schools that were nearby, only Harris (0.01685) and Edison (0.00187) had a significant impact on the price of the home, with Edison being more significant than Harris. This could mean that these two are prestigious schools, or that they are located near more expensive houses, though it does not imply that these schools are necessarily the cause of the price raise.

After looking at the p-values, I plotted the residuals in order to see if there is any problem regarding the data with the lower $p$-values. The residuals for the age and the garage appear to be spread well enough that there may just not be a correlation, but the residuals associated with the dummy variables seemed to be a bit heavy on one side and not well spread on the other, suggesting that they may not contain an accurate depiction of the dummy variables. An example of the problems with the residuals of the dummy variables can be seen with the Parker residuals graph shown below.


When I was done looking at all of the p-values separately, I looked at the graph of the residuals of the whole equation. The mean of the residuals had a fairly constant mean of zero, but the variance appears to have a funnel or cone shape that gets much larger to the right of the graph. The qq line kept showed that the residuals were normally distributed. The graphs of the residuals and the qq line are shown below.



After looking at both the p -values and the residuals of the variables, a new model was used based on the factors with the lower $p$-values. The factors in this new model and their original $p$-values are Size (0.049), Lot (0.00518), Bath (0.01961), Bed (0.01961), Bathbed (0.0338), Agesq (0.01905), Active (0.01614), Edison (0.00187), and Harris (0.01685). These factors together give a p-value of 8.821( $\left.\mathrm{e}^{\wedge} 9\right)$. Although there is not much of a change in the overall $p$-value, it seems that the individual $p$-values are low enough to support the idea that these values fit the line the best. It may also signify that there are many values outside of this dataset that have not yet been explored. The new $p$-values for each of the factors are Size ( 0.004664 ), Lot ( 0.002776 ), Bath ( 0.061327 ), Bed ( 0.005313 ), Bathbed ( 0.026411 ), Agesq (0.091283), Active (0.005476), Edison (0.000315), and Harris (0.001001).

After the $p$-values of the new model were calculated, the new interaction plot was analyzed. The change in the factors gave more residuals that had a mean that appeared to equal zero. The variance was not well spread, though there appeared to be almost no interactions between factors excluding the Bathbed, Bed, and Bath factors' interaction. With many of the factors, there is an issue with constant variance, though this issue can be argued for each factor, and therefore these variables can be left in the formula. When all of the residuals are taken together, the variance appears to have improved due to the fact that the cone shape of the residuals are not as drastic. A model of the residuals when taken together is shown below, along with a model of the qq line, which showed that the residuals stayed normally distributed.


The rise in the p-value of Agesq suggests that a closer inspection of the factor is appropriate. Looking at the residuals of Agesq, five points can be seen that appear to be off to the side. These can skew the mean of the residuals and influence the values of the factor, which shows that it may not be a good depiction of the actual value of age squared. After that, there is the fact that there is a cluster of data points near the y axis, which does not appear to be significant since they are centered somewhere around 0 . Even though these problems exist, the data is fairly distributed and there appears to be a constant variance, leading me to believe that the rest of the data points account and make up for this deviation of the later residuals. The graph of the residuals of Agesq is shown below.


After analyzing the residuals associated with the Agesq factor, I analyzed the residuals associated with the other factors. For the dummy variables, there appears to be a larger spread and more variables on the right side, suggesting that they are better indicators for the model than the other dummy variables mentioned earlier. Some of the residuals for the other factors are questionable, such as the size, in which there is a slight chance that there may not be a constant variance. Most of the variance problems are caused by a single data point, however, so it still seems that this factor can be used as a fair predictor of price.

After the first two equations had been analyzed, I made one last equation using stepwise regression. Using the Aikaike Information Criterion, I obtained a new equation where the factors Size, Lot, Bath, Bed, Bathbed, Agesq, Garage, Active, Edison, and Harris were used. The interesting part of this equation is that all of the same factors are used in the second equation, except for garage, which must have had a higher p -value when only the other factors were used. This could be due to some multicollinearity issues that can be resolved when the other factors are taken out. Taken together, the Aikaike Information Criterion (AIC) of the first equation used is 800.4267 , the AIC of the second equation used is 798.6595 , and the AIC of the third equation is 795.6837 . This shows a slight improvement from the original equation, though not by much.

Looking at the residuals, the last equation fixes many of the issues associated with the first two models. The mean stays at about zero, the residuals follow the normality assumption as shown by the qq
plot, and there is almost a constant variance. There is still a slight cone shape, but, for the most part, that has been fixed, making it more usable. The residual plot and qq line are shown below.


After analyzing all of the residuals together, I analyzed them separately.

While there were still some issues associated with the spread of the residuals and some clustering problems, especially with Agesq which seemed to be having the same issues that it had earlier, the overall spread of these factors had improved. This may have occurred because stepwise regression took out the factors that were having the most issues, thus insuring that the factors that predicted the price best were used. In many of the residuals, there was not a constant variance, yet all other assumptions appeared to have been met. For instance, many had a mean that was equal to zero, though, to some extent, it could be argued that knowing one residual may help a person to better predict the location of another residual due to the shape of them. An example of one of the residual plots is shown below, in which the residuals of the factor Garage are shown.


The last procedure that was done was to make one last regression line model in which the factors were squared in order to see if there was any interaction between or within the factors during the analysis. Unfortunately, this yielded no results, showing that the optimal equation that can be used is equation three. None of the equations could account for most of the prices, showing that there may be other factors at work when deciding the price of the home. Many of these factors do still appear to have an effect, however. More studies can be done in order to see what other equation can influence the price of a home, though this is the best that can be done with the information provided.

## Appendix

## Code

```
attach(homesr)
pairs(~ ., data=homesr)
homes.model1 <- Im(price~., data = homesr)
summary(homes.model1)
plot (homesr$age, homesr$price)
plot (homesr$agesq, homesr$price)
res <- residuals(homes.model1)
plot (size, res)
plot (lot, res)
plot (bath,res)
plot (bed,res)
plot (bathbed,res)
plot (age,res)
plot (agesq,res)
plot (garage,res)
plot (active,res)
plot (edison,res)
plot (harris,res)
plot (adams,res)
plot (crest,res)
plot (parker,res)
qqnorm(res)
qqline(res)
homes.model2 <- Im(price~size+lot+bath+bed+bathbed+agesq+active+edison+harris, data=homesr)
summary(homes.model2)
pairs(~size+lot+bath+bed+bathbed+agesq+active+edison+harris)
plot(homes.model1)
plot(homes.model2)
step (homes.model1)
homes.model3 <- step (homes.model1)
summary(homes.model3)
plot(homes.model3)
pairs(~size+lot+bath+bed+bathbed+agesq+garage+active+edison+harris)
AIC(homes.model1)
AIC(homes.model2)
AIC(homes.model3)
res2 <- residuals(homes.model2)
plot (size, res2)
plot (lot, res2)
plot (bath,res2)
```

```
plot (bed,res2)
plot (bathbed,res2)
plot (age,res2)
plot (agesq,res2)
plot (garage,res2)
plot (active,res2)
plot (edison,res2)
plot (harris,res2)
plot (adams,res2)
plot (crest,res2)
plot (parker,res2)
res3 <- residuals(homes.model3)
plot (size, res3)
plot (lot, res3)
plot (bath,res3)
plot (bed,res3)
plot (bathbed,res3)
plot (age,res3)
plot (agesq,res3)
plot (garage,res3)
plot (active,res3)
plot (edison,res3)
plot (harris,res3)
plot (adams,res3)
plot (crest,res3)
plot (parker,res3)
qqnorm(res2)
qqline(res2)
qqnorm(res3)
qqline(res3)
homes.model4 <-Im(price~.^2,homesr)
summary(homes.model4)
```


## Model Summaries

```
> homes.mode11 <- 1m(price~., data = homesr)
> summary(homes.mode11)
```

Ca11:
1m(formula = price ~ ., data = homesr)
Residuals:

| Min | Madian | 3 Q | Max |  |
| ---: | ---: | ---: | ---: | ---: |
| -83.284 | -22.628 | -0.066 | 27.790 | 111.323 |

Coefficients:

|  | Estimate | Std. Error | t value | $\operatorname{Pr}(>\|t\|)$ |
| :--- | ---: | ---: | ---: | ---: |
| (Intercept) | 337.6628 | 124.9621 | 2.702 | 0.00891 | \%

```
lrrest 
Signif. codes: 0 ‘***' 0.001 ‘**’ 0.01 ‘*’ 0.05 '.' 0.1 ' ' 1
Residual standard error: 42.37 on 61 degrees of freedom
Multiple R-squared: 0.5989, Adjusted R-squared: 0.5068
F-statistic: 6.505 on 14 and 61 DF, p-value: 8.906e-08
```

> homes.mode12 <- 1m(price~size+lot+bath+bed+bathbed+agesq+active+edison+harris, data=
homesr)
> summary (homes.mode12)
Ca11:
1m(formula = price ~ size + lot + bath + bed + bathbed + agesq +
active + edison + harris, data $=$ homesr)
Residuals:

| Min | 10 | Median | 3 Q | Max |
| :--- | ---: | ---: | ---: | ---: |
| -101.578 | -23.721 | 0.133 | 27.577 | 110.508 |

Coefficients:

|  | Estimate | Std. | value |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Intercept) | 319.093 | 108.943 | 2.929 | 0.004664 | ** |
| size | 71.358 | 27.978 | 2.550 | 0.013086 | * |
| 1ot | 10.617 | 3.416 | 3.108 | 0.002776 | ** |
| bath | -82.136 | 43.148 | -1.904 | 0.061327 |  |
| bed | -82.000 | 28.443 | -2.883 | 0.005313 |  |
| bathbed | 27.523 | 12.119 | 2.271 | 0.026411 |  |
| agesq | 1.237 | 0.722 | 1.714 | 0.091283 |  |
| active | 31.853 | 11.090 | 2.872 | 0.005476 |  |
| edison | 62.787 | 16.509 | 3.803 | 0.000315 | ** |
| harris | 51.584 | 14.978 | 3.444 | 0.001001 |  |

Signif. codes: 0 '***' 0.001 ‘**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 43 on 66 degrees of freedom Mu7tiple R-squared: 0.553, Adjusted R-squared: 0.492 F-statistic: 9.072 on 9 and 66 DF, $p$-value: 8.821e-09
summary (homes.mode13)
Ca11:
lm(formula = price $\sim$ size + lot + bath + bed + bathbed + agesq + garage + active + edison + harris, data $=$ homesr)
Residuals:

| Min | 10 | Median | 3 Q | Max |
| ---: | ---: | ---: | ---: | ---: |
| -84.234 | -24.091 | -0.581 | 29.075 | 104.106 |

Coefficients:

|  | Estimate | Std. Error | value | $\operatorname{Pr}(>\|t\|)$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Intercept) | 322.7285 | 106.2569 | 3.037 | 0.003435 | ** |
| size | 59.6082 | 27.8541 | 2.140 | 0.036112 | * |
| lot | 9.3038 | 3.3893 | 2.745 | 0.007815 | ** |
| bath | -92.8312 | 42.3867 | -2.190 | 0.032111 | * |
| bed | -79.6453 | 27.7610 | -2.869 | 0.005549 | ** |
| bathbed | 29.5347 | 11.8574 | 2.491 | 0.015308 | * |
| agesq | 1.4557 | 0.7118 | 2.045 | 0.044878 | * |
| garage | 16.2220 | 7.7353 | 2.097 | 0.039875 | * |
| active | 27.7465 | 10.9913 | 2.524 | 0.014040 | * |
| edison | 61.9646 | 16.1050 | 3.848 | 0.000275 | *** |
| harris | 49.9654 | 14.6272 | 3.416 | 0.001100 | ** |

Signif. codes: 0 ‘***' 0.001 ‘**' 0.01 ‘*’ 0.05 '.’ 0.1 ' ' 1
Residual standard error: 41.93 on 65 degrees of freedom
Multiple R-squared: 0.5813, Adjusted R-squared: 0.5169
F-statistic: 9.025 on 10 and 65 DF, p-value: $4 \mathrm{e}-09$

```
> homes.mode14 <-1m(price~.^2,homesr)
> summary(homes.mode14)
```

Ca11:
1m(formula $=$ price $\sim$.^2, data $=$ homesr)

|  | 2 | 3 | 4 | 5 |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| -9.820e-15 | $7.819 \mathrm{e}-15$ | 8.766e-15 | $5.991 \mathrm{e}-16$ | -3.992e-16 | $5.703 \mathrm{e}-16$ |
|  |  |  | 10 | 11 | 12 |
| -2.241e-15 | -3.763e-14 | 8.227e-15 | $7.178 \mathrm{e}-15$ | -3.850e-14 | $-7.929 \mathrm{e}-15$ |
| 13 | 14 | 15 | 16 | 17 | 18 |
| $5.925 \mathrm{e}-14$ | $-1.218 \mathrm{e}-14$ | -5.978e-16 | $1.083 \mathrm{e}-14$ | $-1.884 e-15$ | 347e-13 |
| 19 | 20 | 21 | 22 | 23 | 24 |
| -4.215e-13 | -6.712e-15 | $2.457 \mathrm{e}-16$ | $-1.004 e-14$ | .711e-14 | $334 \mathrm{e}-14$ |
| 25 | 26 | 27 | 28 | 29 | 30 |
| 5.942e-14 | 1.569e-14 | -2.419e-14 | 1.587e-14 | -4.389e-14 | -1.045e-14 |
| 31 | 32 | 33 | 34 | 35 | 36 |
| $1.733 \mathrm{e}-16$ | -1.649e-16 | 330e-14 | -6.043e-15 | -2.574e-15 | e-15 |
| 37 | 38 | 39 | 40 | 41 | 42 |
| $6.612 \mathrm{e}-15$ | -6.411e-14 | -1.354e-14 | 3.064e-14 | 798e-16 | $1.460 \mathrm{e}-14$ |
| 43 | 44 | 45 | 46 | 47 | 48 |
| -4.235e-14 | -3.183e-15 | 972e-14 | $-1.129 \mathrm{e}-14$ | 139e-14 | -9.345e-14 |
| 49 | 50 | 51 | 52 | 53 | 54 |
| -1.160e-16 | $2.469 \mathrm{e}-16$ | -1.074e-14 | $-1.610 e-14$ | .085e-15 | 085e-15 |
| 55 | 56 | 57 | 58 | 59 | 60 |
| -3.346e-15 | -1.597e-14 | $316 \mathrm{e}-14$ | $1.044 \mathrm{e}-14$ | $625 \mathrm{e}+00$ | $-4.333 e+00$ |
| 61 | 62 | 63 | 64 | 65 | 66 |
| $-5.625 \mathrm{e}+00$ | -3.333e-01 | $4.667 \mathrm{e}+00$ | . 965e-15 | -1.786e-14 | -1.774e-14 |
| 67 | 68 | 69 | 70 | 71 | 72 |
| $3.690 \mathrm{e}-13$ | -5.787e-14 | $-1.958 \mathrm{e}-16$ | -3.632e-14 | $3.233 \mathrm{e}-16$ | 7.426e-16 |
| 73 | 74 | 75 | 76 |  |  |
| $4.438 \mathrm{e}-15$ | 0e-14 | 14 | .086e-16 |  |  |


| C | (33 not defined because of singularities) |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Intercept) | Estimate $30674.55$ | Std. Error 69920.30 | t value | $\operatorname{Pr}(>\mid \mathrm{tl})$ 0.69055 |  |
| size | -42013.34 | 29343.43 | -1.432 | 0.24761 |  |
| lot | -3865.14 | 4652.54 | -0.831 | 0.46705 |  |
| bath | 16682.49 | 36127.34 | 0.462 | 0.67567 |  |
| bed | -9210.35 | 20147.24 | -0.457 | 0.67863 |  |
| bathbed | -507.76 | 9807.71 | -0.052 | 0.96197 |  |
| age | -9862.19 | 4350.87 | -2.267 | 0.10825 |  |
| agesq | 1955.93 | 830.69 | 2.355 | 0.09989 |  |
| garage | 16761.83 | 6086.66 | 2.754 | 0.07051 |  |
| active | -6524.18 | 1680.91 | -3.881 | 0.03030 |  |
| edison | 20869.93 | 7045.49 | 2.962 | 0.05944 |  |
| harris | 154178.77 | 51793.94 | 2.977 | 0.05875 |  |
| adams | 11430.87 | 4831.29 | 2.366 | 0.09886 |  |
| crest | 3630.37 | 6095.82 | 0.596 | 0.59341 |  |
| parker | -1985.29 | 5832.97 | -0.340 | 0.75603 |  |
| size:lot | 1226.20 | 293.55 | 4.177 | 0.02499 | * |
| size:bath | 6197.07 | 14029.61 | 0.442 | 0.68860 |  |
| size:bed | 10011.33 | 8515.16 | 1.176 | 0.32452 |  |
| size:bathbed | -2068.51 | 4061.57 | -0. 509 | 0.64565 |  |
| size:age | -594.70 | 146.81 | -4.051 | 0.02709 | * |
| size:agesq | 184.12 | 27.62 | 6.666 | 0.00688 |  |
| size:garage | 2840.82 | 692.95 | 4.100 | 0.02626 |  |
| size:active | 3127.10 | 785.70 | 3.980 | 0.02838 | * |
| size:edison | -1987. 39 | 410.57 | -4.841 | 0.01682 | * |
| size:harris | -6804.34 | 1895.59 | -3.590 | 0.03703 | * |
| size:adams | -5863.16 | 2515.39 | -2.331 | 0.10206 |  |
| size:crest | 1426.01 | 2165.55 | 0.658 | 0.55724 |  |
| size:parker | -5981.39 | 1683.58 | -3.553 | 0.03802 | * |
| lot:bath | -384.98 | 2310.86 | -0.167 | 0.87828 |  |
| lot:bed | 378.31 | 1433.43 | 0.264 | 0.80893 |  |
| 1ot: bathbed | 18.01 | 707.46 | 0.025 | 0.98129 |  |
| lot:age | 36.92 | 33.22 | 1.111 | 0.34746 |  |
| 1ot:agesq | 17.87 | 24.12 | 0.741 | 0.51243 |  |


| 1ot:garage | 374.14 | 112.38 | 3.329 | 0.04474 | * |
| :---: | :---: | :---: | :---: | :---: | :---: |
| lot:active | -157.72 | 57.10 | -2.762 | 0.07002 |  |
| lot:edison | 910.82 | 265.18 | 3.435 | 0.04139 | * |
| 1ot:harris | 508.00 | 137.70 | 3.689 | 0.03454 | * |
| lot:adams | 281.50 | 126.03 | 2.234 | 0.11162 |  |
| 1ot:crest | 2013.32 | 686.57 | 2.932 | 0.06088 |  |
| 1ot:parker | -143.24 | 131.72 | -1.087 | 0.35640 |  |
| bath:bed |  | NA | NA | NA |  |
| bath: bathbed | -457.60 | 201.29 | -2.273 | 0.10759 |  |
| bath:age | 5107.28 | 2127.53 | 2.401 | 0.09583 |  |
| bath:agesq | -1654.59 | 570.97 | -2.898 | 0.06261 |  |
| bath:garage | -11795.23 | 3622.55 | -3.256 | 0.04727 | * |
| bath:active | -104.81 | 612.53 | -0.171 | 0.87502 |  |
| bath:edison | -5932.15 | 2849.01 | -2.082 | 0.12873 |  |
| bath:harris | -70086.72 | 23770.60 | -2.948 | 0.06010 |  |
| bath:adams | NA | NA | NA | NA |  |
| bath:crest | -6642.94 | 1467.22 | -4.528 | 0.02016 | * |
| bath: parker | 7948.23 | 4124.82 | 1.927 | 0.14962 |  |
| bed:bathbed | -97.59 | 60.03 | -1.626 | 0.20248 |  |
| bed:age | 2299.10 | 1091.01 | 2.107 | 0.12569 |  |
| bed: agesq | -680.43 | 257.98 | -2.638 | 0.07782 |  |
| bed:garage | -5979.90 | 1878.84 | -3.183 | 0.04999 | * |
| bed: active | 273.35 | 253.91 | 1.077 | 0.36055 |  |
| bed:edison | -6674.89 | 2498.34 | -2.672 | 0.07558 |  |
| bed:harris | -46978.71 | 16484.12 | -2.850 | 0.06511 |  |
| bed: adams | NA | NA | NA | NA |  |
| bed:crest | -887.93 | 594.57 | -1.493 | 0.23217 |  |
| bed:parker | 4753.38 | 2466.37 | 1.927 | 0.14956 |  |
| bathbed:age | -1263.95 | 562.50 | -2.247 | 0.11024 |  |
| bathbed:agesq | 465.56 | 166.24 | 2.800 | 0.06783 |  |
| bathbed:garage | 3078.07 | 938.80 | 3.279 | 0.04647 | * |
| bathbed:active | 48.62 | 143.68 | 0.338 | 0.75739 |  |
| bathbed:edison | 1788.62 | 939.69 | 1.903 | 0.15312 |  |
| bathbed:harris | 23010.28 | 7943.94 | 2.897 | 0.06268 |  |
| bathbed:adams | NA | NA | NA | NA |  |
| bathbed:crest | 187.81 | 370.05 | 0.508 | 0.64674 |  |
| bathbed: parker | -2651.38 | 1312.46 | -2.020 | 0.13664 |  |
| age: agesq | 73.89 | 22.30 | 3.313 | 0.04530 | * |
| age:garage | 430.62 | 114.42 | 3.764 | 0.03281 | * |
| age:active | -261.29 | 68.35 | -3.823 | 0.03151 | * |
| age:edison | -332.89 | 95.12 | -3.500 | 0.03949 | * |
| age:harris | 40.40 | 46.99 | 0.860 | 0.45319 |  |
| age:adams | NA | NA | NA | NA |  |
| age:crest | NA | NA | NA | NA |  |
| age:parker | 51.72 | 54.63 | 0.947 | 0.41360 |  |
| agesq:garage | 89.36 | 35.23 | 2.537 | 0.08493 |  |
| agesq:active | NA | NA | NA | NA |  |
| agesq:edison | NA | NA | NA | NA |  |
| agesq:harris | NA | NA | NA | NA |  |
| agesq:adams | NA | NA | NA | NA |  |
| agesq:crest | NA | NA | NA | NA |  |
| agesq:parker | NA | NA | NA | NA |  |
| garage:active | NA | NA | NA | NA |  |
| garage:edison | NA | NA | NA | NA |  |
| garage:harris | NA | NA | NA | NA |  |
| garage:adams | NA | NA | NA | NA |  |
| garage:crest | NA | NA | NA | NA |  |
| garage:parker | NA | NA | NA | NA |  |
| active:edison | NA | NA | NA | NA |  |
| active:harris | NA | NA | NA | NA |  |
| active:adams | NA | NA | NA | NA |  |
| active:crest | NA | NA | NA | NA |  |
| active:parker | NA | NA | NA | NA |  |
| edison:harris | NA | NA | NA | NA |  |
| edison:adams | NA | NA | NA | NA |  |
| edison:crest | NA | NA | NA | NA |  |
| edison:parker | NA | NA | NA | NA |  |
| harris:adams | NA | NA | NA | NA |  |
| harris:crest | NA | NA | NA | NA |  |
| harris:parker | NA | NA | NA | NA |  |
| adams:crest | NA | NA | NA | NA |  |
| adams:parker | NA | NA | NA | NA |  |


| crest:parker | NA | NA | NA | NA |
| :---: | :---: | :---: | :---: | :---: |
| Signif. codes: |  |  |  |  |

Residual standard error: 5.886 on 3 degrees of freedom Multiple R-squared: 0.9996, Adjusted R-squared: 0.9905 F-statistic: 109.4 on 72 and 3 DF, p-value: 0.00121

## Stepwise Regression



```
- bed 1 15696.2 126593 587.77
- harris 1 16508.4 127406 588.25
- edison 1 24361.9 135259 592.80
Step: AIC=578.73
price ~ size + lot + bath + bed + bathbed + age + agesq + garage +
    active + edison + harris
\begin{tabular}{lrrrrr} 
& Df & Sum of Sq & RSS & AIC \\
- age & 1 & 1905.4 & 114301 & 578.01 \\
<none> & & & 112395 & 578.73 \\
- garage & 1 & 4403.5 & 116799 & 579.65 \\
- size & 1 & 7219.8 & 119615 & 581.46 \\
- agesq & 1 & 8799.0 & 121194 & 582.46 \\
- bath & 1 & 9294.7 & 121690 & 582.77 \\
- active & 1 & 10938.1 & 123334 & 583.79 \\
- bathbed & 1 & 11495.9 & 123891 & 584.13 \\
- bed & 1 & 14198.6 & 126594 & 585.77 \\
- lot & 1 & 14612.8 & 127008 & 586.02 \\
- harris & 1 & 17810.5 & 130206 & 587.91 \\
- edison & 1 & 27910.3 & 140306 & 593.58
\end{tabular}
Step: AIC=578.01
price ~ size + lot + bath + bed + bathbed + agesq + garage +
        active + edison + harris
<none> Df Sum of Sq RSS R AIC
- agesq 1 7356.1 121657 580.75
- garage 1 7733.8 122035 580.98
- size 1 8053.2 122354 581.18
- bath 1 8434.6 122735 581.42
- bathbed 1 10910.0 125211 582.93
- active 1 11206.1 125507 583.11
- lot 1}1313250.4 127551 584.3
- bed 1 14473.9 128775 585.07
- harris 1 20518.9 134820 588.55
- edison 1 26031.6 140332 591.60
```

